
Common Concentration Inequalities

Beier Zhu

Markov's inequality: basis for the rest inequalities

Theorem 1. (Markov's inequality). If X is a non-negative r.v. and $\mu = \mathbb{E}[X]$, then $\forall t > 0$:

$$\mathbb{P}[X \geq t] \leq \frac{\mu}{t} \quad (1)$$

Polynomial tail bounds $\mathcal{O}(t^{-k})$

Derivation insight: Markov's inequality + bounded k -th central moment

Theorem 2. (Polynomial variant of markov's inequality). If X is a r.v. with mean μ and finite k -th central moment $\mathbb{E}[|X - \mu|^k]$, then $\forall t > 0$,

$$\mathbb{P}[|X - \mu| \geq t] \leq \frac{\mathbb{E}[|X - \mu|^k]}{t^k} \quad (2)$$

Specifically, when $k = 2$ and denote the variance as σ^2 , we have:

Theorem 3. (Chebyshev's inequality). If X is a r.v. with mean μ and variance σ^2 , then $\forall t > 0$

$$\mathbb{P}[|X - \mu| \geq t] \leq \frac{\sigma^2}{t^2} \quad (3)$$

Theorem 4. (Chebyshev's inequality for the sample mean). Let X_1, \dots, X_n be i.i.d with mean μ and variance σ^2 . Define $\nu = \frac{1}{n} \sum_{i=1}^n X_i$. Then, $\forall t > 0$

$$\mathbb{P}[|\nu - \mu| \geq t] \leq \frac{\sigma^2}{nt^2} \quad (4)$$

Theorem 5. (Cantelli's inequality, a.k.a. one-sided chebyshev's inequality). If X is a r.v. with mean μ and variance σ^2 , then $\forall t > 0$

$$\mathbb{P}[X - \mu \geq t] \leq \frac{\sigma^2}{\sigma^2 + t^2}, \quad \mathbb{P}[X - \mu \leq -t] \leq \frac{\sigma^2}{\sigma^2 + t^2} \quad (5)$$

Exponential tail bounds $\mathcal{O}(\exp\{-t^2\})$

Derivation insight: Markov inequality + bounded moment generating function

Theorem 6. (Chernoff's inequality). For a r.v. X with finite moment generating function $M_X(\lambda)$, we have

$$\mathbb{P}[X - \mu \geq t] \leq \inf_{\lambda \geq 0} M_X(\lambda) \exp\{-\lambda(t + \mu)\} \quad (6)$$

Theorem 7. (Tail bound for sub-Gaussian r.v.) If a r.v. X with finite mean μ is σ -sub-Gaussian, then $\forall t > 0$

$$\mathbb{P}[|X - \mu| \geq t] \leq 2 \exp\left\{-\frac{t^2}{2\sigma^2}\right\} \quad (7)$$

Theorem 8. (Hoeffding's inequality). Let X_1, \dots, X_n be independent real-valued r.v.s drawn from some distribution, such that $a_i \leq X_i \leq b_i$ almost surely. Define $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, and let

$\mu = \mathbb{E}[\bar{X}]$. Then $\forall t > 0$,

$$\mathbb{P}[|\bar{X} - \mu| \geq t] \leq 2 \exp\left\{-\frac{2n^2 t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right\} \quad (8)$$

Corollary 1. (Hoeffding's inequality for the sample mean). Let X_1, \dots, X_n be i.i.d r.v.s with $a \leq X_i \leq b$. Then $\forall t > 0$,

$$\mathbb{P}[|\bar{X} - \mu| \geq t] \leq 2 \exp\left\{-\frac{2nt^2}{(b-a)^2}\right\} \quad (9)$$

Specifically, when $n = 1$: If X is a r.v. with $a \leq X \leq b$. Then $\forall t > 0$,

$$\mathbb{P}[|X - \mu| \geq t] \leq 2 \exp\left\{-\frac{2t^2}{(b-a)^2}\right\} \quad (10)$$

Motivation: Sub-Gaussian tails scale with the variance, while Hoeffding's bound depends only on the range of the variables. For bounded variables with **small variance**, this suggests we can obtain sharper concentration bounds than Hoeffding's bound.

Theorem 9. (Bernstein's inequality). Let X_1, \dots, X_n be i.i.d r.v.s with $a \leq X_i \leq b$ and $\mathbb{V}[X] = \sigma^2$. Then $\forall t > 0$

$$\mathbb{P}[|\bar{X} - \mu| \geq t] \leq 2 \exp\left\{-\frac{nt^2}{2(\sigma^2 + (b-a)t)}\right\} \quad (11)$$

Theorem 10. (χ^2 tail bound). Suppose that $X_1, \dots, X_n \sim \mathcal{N}(0, 1)$, then $\forall t \in (0, 1)$:

$$\mathbb{P}\left[\left|\frac{1}{n} \sum_{i=1}^n X_i^2 - 1\right| \geq t\right] \leq 2 \exp\left\{-\frac{nt^2}{8}\right\} \quad (12)$$

Concentrations of functions of r.v.s

Motivation: So far we have focused on the contraction of averages. A natural question is whether other functions of i.i.d r.v.s also show exponential concentration.

Theorem 11. (McDiarmid's inequality). Suppose $f : \mathbb{R}^n \rightarrow \mathbb{R}$ satisfies the **bounded difference condition**: there exist constants $c_1, \dots, c_n \in \mathbb{R}$ such that for all real numbers x_1, \dots, x_n and x'_i ,

$$|f(x_1, \dots, x_n) - f(x_1, \dots, x_{i-1}, x'_i, x_{i+1}, \dots, x_n)| \leq c_i \quad (13)$$

Intuitively, Eq. (13) states that f is not overly sensitive to arbitrary changes in a single coordinates. Then for any independent random variables X_1, \dots, X_n ,

$$\mathbb{P}\{f(X_1, \dots, X_n) - \mathbb{E}[f(X_1, \dots, X_n)] \geq t\} \leq \exp\left[-\frac{2t^2}{\sum_{i=1}^n c_i^2}\right]. \quad (14)$$

Moreover, $f(X_1, \dots, X_n)$ is $\mathcal{O}(\sqrt{\sum_{i=1}^n c_i^2})$ -sub-Gaussian.

Motivation: The bounded difference in McDiarmid's inequality is often satisfied by bounded r.v.s or a bounded function. To get similar concentration inequalities for unbounded r.v.s like Gaussians, we need some other special conditions.

Theorem 12. Suppose $f : \mathbb{R}^n \rightarrow \mathbb{R}$ is L -Lipschitz w.r.t ℓ_2 -norm, and X_1, \dots, X_n drawn i.i.d. from $\mathcal{N}(0, 1)$. Then, $\forall t \in \mathbb{R}$,

$$\mathbb{P}[|f(X_1, \dots, X_n) - \mathbb{E}[f(X_1, \dots, X_n)]| \geq t] \leq 2 \exp\left(-\frac{t^2}{2L^2}\right) \quad (15)$$

1 Proofs

1.1 Proof of Theorem 1: Markov's inequality

$$t\mathbb{P}[X \geq t] = t \int_t^\infty \mathbb{P}(x)dx = \int_t^\infty t\mathbb{P}(x)dx \quad (16)$$

$$\leq \int_t^\infty x\mathbb{P}(x)dx \leq \int_0^\infty x\mathbb{P}(x)dx \quad (17)$$

$$= \mathbb{E}[X] \quad (18)$$

1.2 Proof of Theorem 2: Polynomial variant of Markov's inequality

$$\mathbb{P}[|X - \mu| \geq t] = \mathbb{P}[|X - \mu|^k \geq t^k] \quad (19)$$

$$\leq \frac{\mathbb{E}[|X - \mu|^k]}{t^k} \quad [\text{Markov's Inequality}] \quad (20)$$

1.3 Proof of Theorem 4: Chebyshev's inequality for the sample mean

$$\mathbb{V}[\nu] = \frac{1}{n^2} \sum_{i=1}^n \mathbb{V}[X_i] = \frac{\sigma^2}{n}. \quad (21)$$

According to Theorem 3:

$$\mathbb{P}[|\nu - \mu| \geq t] \leq \frac{\mathbb{V}[\nu]}{t^2} = \frac{\sigma^2}{nt^2} \quad (22)$$

1.4 Proof of Theorem 5: Cantelli's inequality

Let $Y = X - \mu$, then $\mathbb{E}[Y] = 0$ and $\mathbb{V}[Y] = \sigma^2$. For any λ s.t. $t + \lambda > 0$, we have:

$$\mathbb{P}[Y \geq t] = \mathbb{P}[Y + \lambda \geq t + \lambda] \quad (23)$$

$$= \mathbb{P}\left[\frac{Y + \lambda}{t + \lambda} \geq 1\right] \quad [t + \lambda > 0] \quad (24)$$

$$\leq \mathbb{P}\left[\left(\frac{Y + \lambda}{t + \lambda}\right)^2 \geq 1\right] \quad (25)$$

$$\leq \mathbb{E}\left[\left(\frac{Y + \lambda}{t + \lambda}\right)^2\right] \quad [\text{Markov's inequality}] \quad (26)$$

$$= \frac{\sigma^2 + t^2}{(\lambda + t)^2} \quad (27)$$

We pick λ to minimize the R.H.S, which is $\lambda = \frac{\sigma^2}{t} > 0$. That proves the theorem.

1.5 Proof of Theorem 6: Chernoff's inequality

Define $\mu = \mathbb{E}[X]$. For any $\lambda > 0$, we have

$$\mathbb{P}[X - \mu \geq t] = \mathbb{P}[X \geq \mu + t] \quad (28)$$

$$= \mathbb{P}[\exp\{\lambda X\} \geq \exp\{\lambda(\mu + t)\}] \quad (29)$$

$$\leq \exp\{-\lambda(\mu + t)\} \mathbb{E}[\exp\{\lambda X\}] \quad [\text{Markov's inequality}] \quad (30)$$

$$= \exp\{-\lambda(\mu + t)\} M_X(\lambda) \quad (31)$$

Now λ is a parameter we can choose to get a tight upper bound.

1.6 Proofs for Theorem 7 and 8: tail bound for sub-Gaussian r.v. and Hoeffding's inequalities

See last document.